

# Multitype branching processes conditioned on very late extinction. An example in epidemiology.

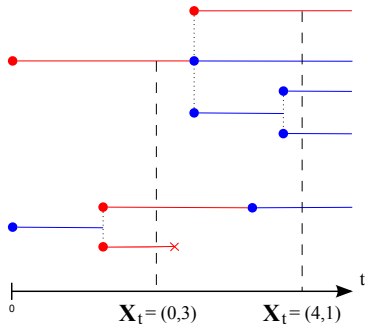
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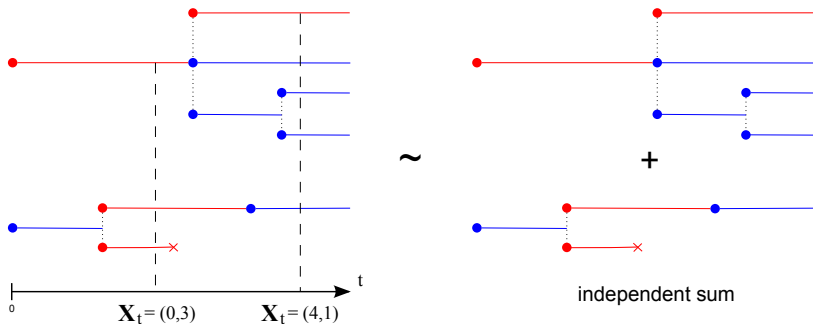
# Multitype branching process

- $\mathbb{N}^d$ -valued Markov process (here  $d = 2$ )
- offspring distribution  $(p_1(\mathbf{k}))_{\mathbf{k} \in \mathbb{N}^d}, (p_2(\mathbf{k}))_{\mathbf{k} \in \mathbb{N}^d}$
- exponentially distributed lifetime with parameter  $\alpha_1, \alpha_2$



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branching property

# Extinction of a multitype BP

mean matrix  $\mathbf{M}$  with entries

$m_{ij} :=$  mean number of offsprings of type  $j$  for a particle of type  $i$

Perron-Frobenius' theorem  $\Rightarrow$  if  $\mathbf{M}$  is **irreducible** then the matrix  $\text{diag}(\alpha_1, \dots, \alpha_d)(\mathbf{M} - \mathbf{I})$  has a unique maximal real eigenvalue  $\rho$  and a positive normalized right eigenvector  $\xi$

## Theorem

$\rho \leq 0 \iff$  *the process dies out almost surely.*

- $\rho < 0$ : **subcritical** process
- $\rho = 0$ : **critical** process
- $\rho > 0$ : **supercritical** process

If  $d = 1$ , then  $\rho \leq 0$  is equivalent to  $m \leq 1$ !

## Conditioning on very late extinction

Irreducible branching process with  $\mathbb{P}$ ,  $\mathbf{M}$ ,  $\rho$ ,  $\xi$

- We want to work with a process which dies out almost surely:
  - critical or subcritical process
  - supercritical process with positive risk of extinction  $\mathbf{q}$ , *conditioned on extinction*

Proposition (Jagers, Lagerås, 2008)

*A supercritical BP conditioned on extinction is a subcritical BP.*

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- We condition this "mortal" BP on very late extinction:

$$\forall t \geq 0, \forall B \in \mathcal{F}_t, \quad \mathbb{P}^*(B) := \lim_{\theta \rightarrow \infty} \mathbb{P} \left( B \mid \mathbf{X}_{t+\theta} \neq \mathbf{0}, \lim_{s \rightarrow \infty} \mathbf{X}_s = \mathbf{0} \right)$$

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Proposition (S. P.)

The conditioned law  $\mathbb{P}^*$  is a  $h$ -transform of the unconditioned law  $\mathbb{P}$

$$d\mathbb{P}_{\mathbf{x}}^*|_{\mathcal{F}_t} = e^{-\tilde{\rho}t} \frac{\mathbf{q}^{\mathbf{x}_t}}{\mathbf{q}^{\mathbf{x}}} \frac{\mathbf{X}_t \cdot \tilde{\xi}}{\mathbf{x} \cdot \tilde{\xi}} d\mathbb{P}_{\mathbf{x}}|_{\mathcal{F}_t}, \quad \mathbf{x} \in \mathbb{N}^d, \quad \mathbf{x} \neq \mathbf{0}.$$



## Conditioning on very late extinction

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In particular, if  $\mathbb{P}$  is (sub)critical:  $d\mathbb{P}_{\mathbf{x}}^*|_{\mathcal{F}_t} = e^{-\rho t} \frac{\mathbf{X}_t \cdot \xi}{\mathbf{x} \cdot \xi} d\mathbb{P}_{\mathbf{x}}|_{\mathcal{F}_t}.$

## Interpretation of the conditioned process

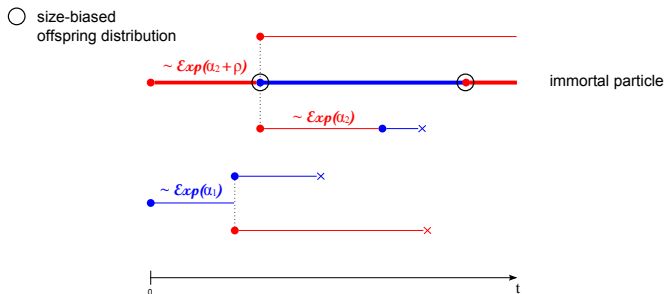
- Conditioning on *extinction* preserves the branching property  
→ modifies the offspring distribution of the BP
- Conditioning on *very late extinction* does not!  
→ adds an external structure to the unconditioned BP

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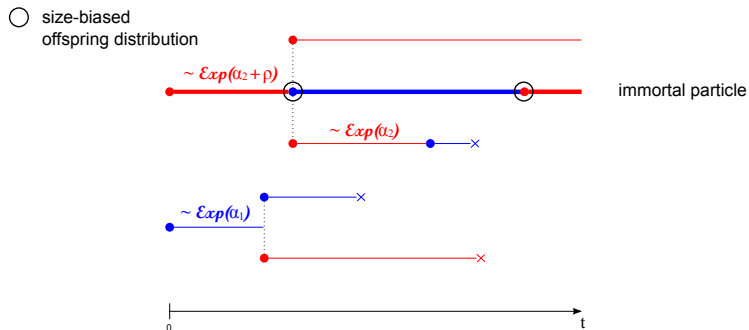
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Proposition (S. P.)

*conditioned process*  $\sim$  *unconditioned process* + *immortal particle*



# Interpretation of the conditioned process



- immortal particle remains of type  $i$  for a time  $\sim \text{Exp}(\alpha_i + \rho)$
- produces  $\mathbf{k}$  offsprings according to the *size-biased* distribution

$$q_i(\mathbf{k}) := \frac{\alpha_i}{(\alpha_i + \rho)\xi_i} \mathbf{k} \cdot \xi p_i(\mathbf{k})$$

- mutates to type  $j$  with probability  $\frac{k_j \xi_j}{\mathbf{k} \cdot \xi}$

## Investigation of other limits

- A rescaled BP converges to the multitype Feller diffusion solution of

$$dX_{t,i} = \sigma_i \sqrt{X_{t,i}} dB_{t,i} + \sum_{j=1}^d c_{ji} X_{t,j} dt, \quad i = 1 \dots d.$$

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- What about the *conditioned* BP? Does it converge to the *conditioned* Feller process solution of

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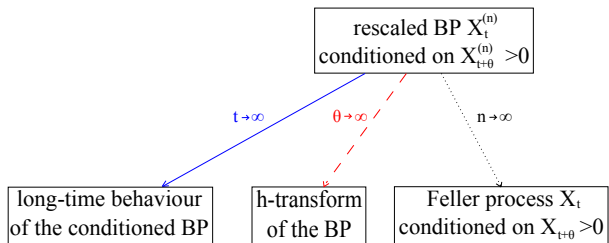
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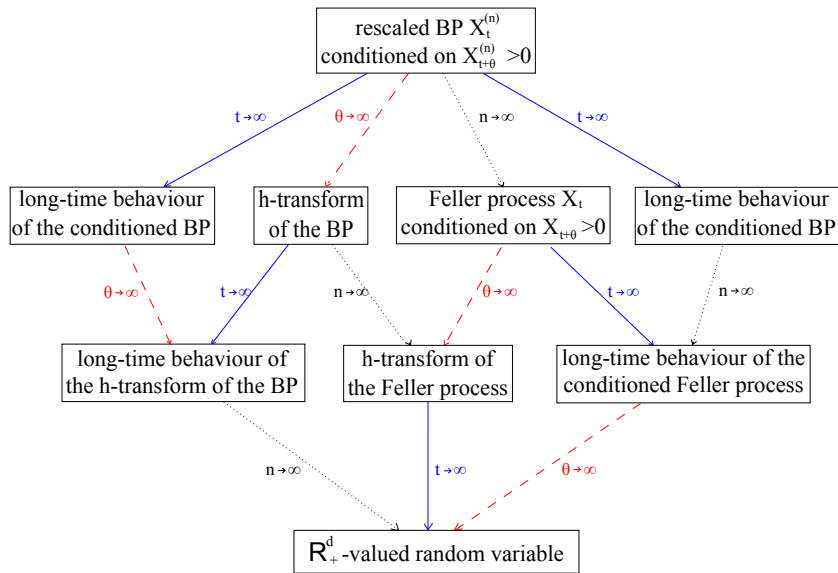
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- What about the long-time behavior of the conditioned processes?
- Three kinds of limits for  $\mathbb{P}^{(n)}(\mathbf{X}_t \in \cdot \mid \mathbf{X}_{t+\theta} \neq \mathbf{0}, \lim_{s \rightarrow \infty} \mathbf{X}_s = \mathbf{0})$ 
  - scaling limit  $n \rightarrow \infty$
  - conditioning on very late extinction  $\theta \rightarrow \infty$
  - asymptotic behaviour  $t \rightarrow \infty$







## Conditioning and rescaling

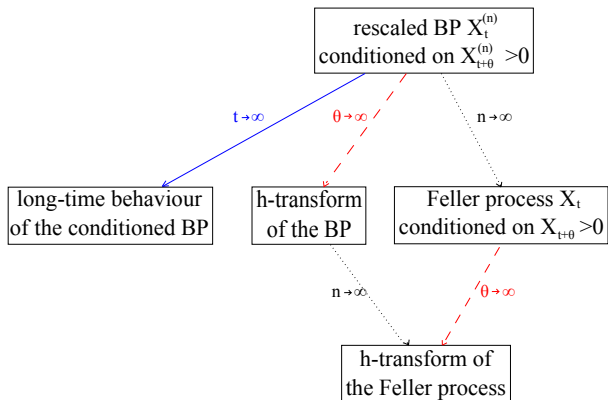
### Theorem (S. P.)

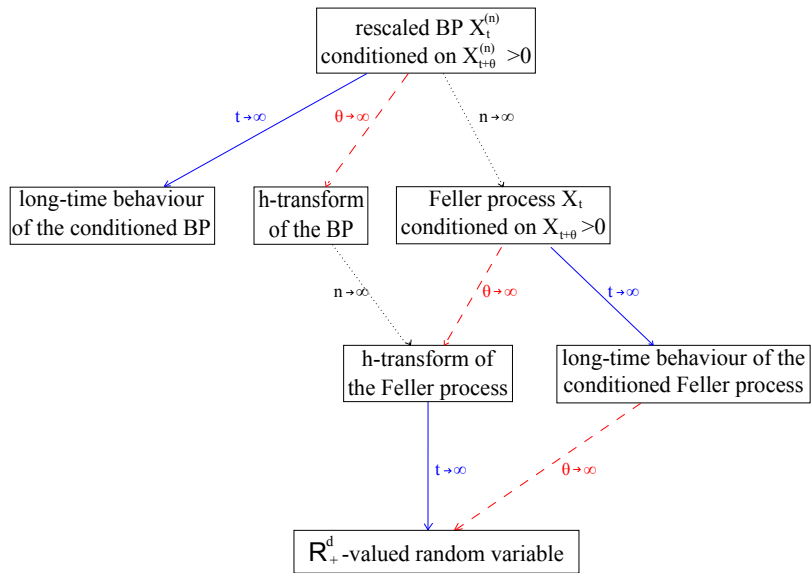
Let  $\mathbf{C}$  be an irreducible mutation matrix. Assume that the offspring distribution  $p_i^{(n)}(\mathbf{j})$  of the rescaled BP satisfies

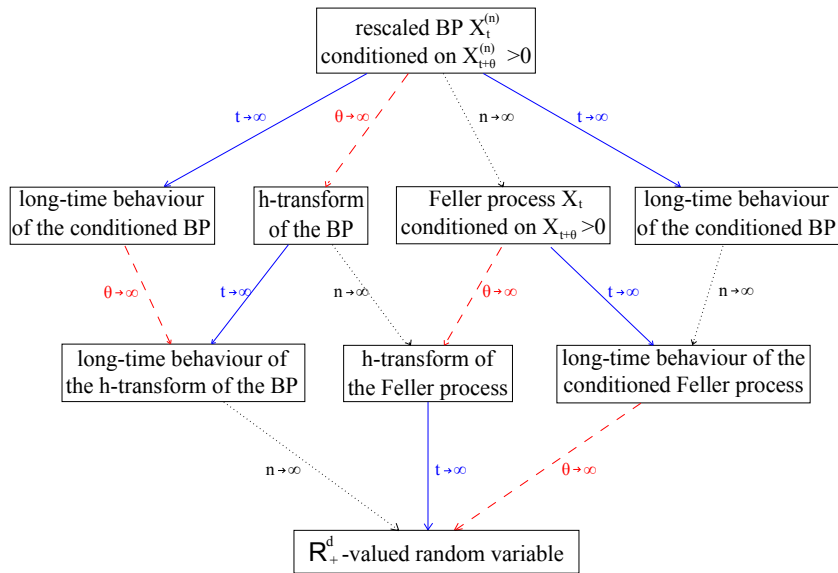
- mean matrix  $\mathbf{M}^{(n)} = \mathbf{I} + \frac{1}{n}\mathbf{C} + o(\frac{1}{n})$
- second order moments  $\sum_{\mathbf{j} \in \mathbb{N}^d} p_i^{(n)}(\mathbf{j})(j_i - 1)^2 = \sigma_i^2 + o(1)$
- $\lim_{N \rightarrow \infty} \sup_{n \in \mathbb{N}} \sum_{\|\mathbf{j}\| > N} \|\mathbf{j}\|^2 p_i^{(n)}(\mathbf{j}) = 0$
- $\sup_{n \in \mathbb{N}} \sum_{\mathbf{j} \in \mathbb{N}^d} p_i^{(n)}(\mathbf{j}) j_i^2 j_k < \infty$

Then, provided the weak convergence of the initial distributions, the following diagram is commutative.

$$\begin{array}{ccc}
 \text{rescaled BP} & \dashrightarrow & \text{conditioned rescaled BP} \\
 \Downarrow & & \Downarrow \\
 \text{Feller diffusion} & \dashrightarrow & \text{conditioned Feller diffusion}
 \end{array}$$

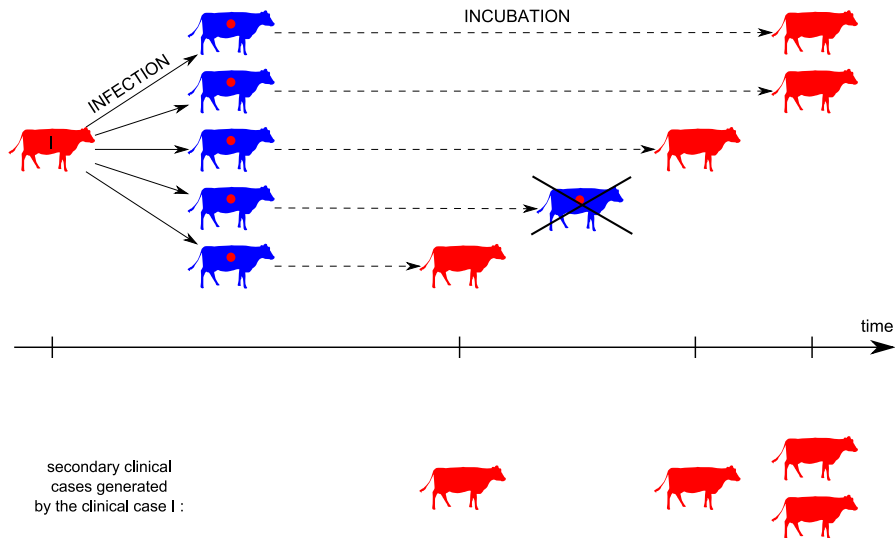




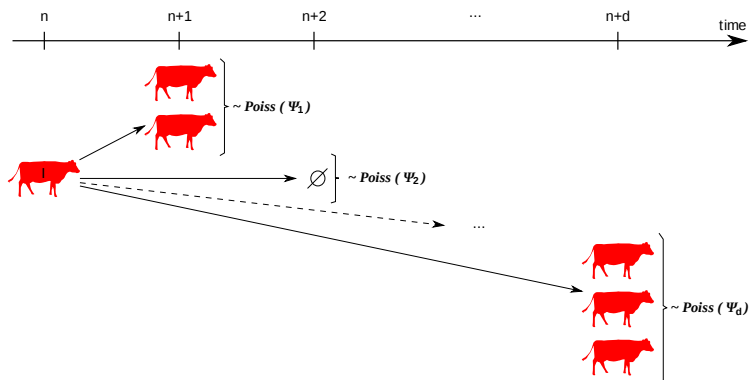


# Application to the BSE epidemic

# The model



# The model



$$\Psi_i = \frac{P_{inc}(i)}{\sum_{k=1}^{d+1} P_{surv}(k)} \left( \theta_{env} \sum_{k=i+1}^{d+1} P_{surv}(k) + p_{mat} P_{surv}(i+1) \right), \quad i = 1 \dots d.$$

$\theta_{env}$  = environmental infection parameter (mean number of newly infected animals via the environment, per infective and per year)



# The model

$X_n$  = number of clinical cases at time  $n$

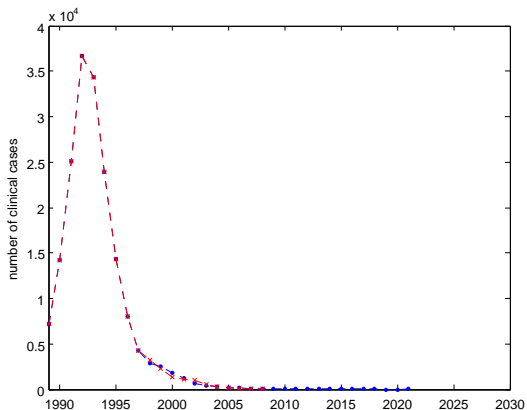
Markovian process of order  $d$ , with transition law

$$X_n | \mathcal{F}_{n-1} \sim \text{Poiss} \left( \sum_{i=1}^d X_{n-i} \Psi_i \right)$$

or Markovian  $d$ -types branching process (types  $\leftrightarrow$  memory), with mean matrix

$$\mathbf{M} = \begin{pmatrix} \Psi_1 & 1 & 0 & \dots & 0 \\ \Psi_2 & 0 & 1 & \dots & 0 \\ \vdots & \vdots & & \ddots & \vdots \\ \Psi_{d-1} & 0 & \dots & \dots & 1 \\ \Psi_d & 0 & \dots & \dots & 0 \end{pmatrix}$$

# Extinction of the epidemic



## Bayesian estimation:

Extinction of the epidemic  $\sim$  2022

Extinction after 2030 with probability 2.5%

## Conditioning on very late extinction

Aim = study of the trajectories with late extinction:

$$\mathbb{P}(\mathbf{X}_n = . | \mathbf{X}_{n+k} \neq \mathbf{0}), \quad k \text{ very large.}$$

Approximation by

$$\mathbb{P}(\mathbf{X}_n^* = .) := \lim_{k \rightarrow \infty} \mathbb{P}(\mathbf{X}_n = . | \mathbf{X}_{n+k} \neq \mathbf{0}).$$

Proposition (S. P.)

$$X_n^* | \mathcal{F}_{n-1}^* \sim \text{Poiss} \left( \sum_{k=1}^d X_{n-k}^* \psi_k \right) + \mathcal{B} \left( \frac{\sum_{i=1}^d X_{n-i}^* \psi_i \xi_1}{\sum_{k=1}^d X_{n-k}^* \psi_k \xi_k + X_{n-1}^* \xi_2 + \dots + X_{n-(d-1)}^* \xi_d} \right)$$

## Estimation

Estimation of the environmental infection parameter  $\theta_{env}$

$$\Psi_i = a_i \theta_{env} + b_i.$$

**Bayesian estimation:**  $\tilde{\theta} = 2.43$

**Conditional Least Squares Estimation:**

"Classical" CLSE estimation

$$\hat{\theta}_{|\mathbf{x}_0|,n} = \arg \min_{\theta} \sum_{k=1}^n \frac{[x_k - \mathbb{E}_{\theta}(X_k | \mathbf{X}_{k-1} = \mathbf{x}_{k-1})]^2}{\sum_{i=1}^d a_i x_{k-i}}$$

"Paranoiac" CLSE estimation

$$\hat{\theta}_{|\mathbf{x}_0|,n}^* = \arg \min_{\theta} \sum_{k=1}^n \frac{[x_k - \mathbb{E}_{\theta}(X_k^* | \mathbf{X}_{k-1}^* = \mathbf{x}_{k-1})]^2}{\sum_{i=1}^d a_i x_{k-i}}$$

# Classical estimation

$$\hat{\theta}_{|\mathbf{x}_0|,n} = \arg \min_{\theta} \sum_{k=1}^n \frac{[X_k - \mathbb{E}_{\theta}(X_k | \mathbf{X}_{k-1} = \mathbf{x}_{k-1})]^2}{\sum_{i=1}^d a_i X_{k-i}}$$

- explicit form of  $\hat{\theta}_{|\mathbf{x}_0|,n}$
- no asymptotic property as  $n \rightarrow \infty$
- consistency and asymptotic behaviour as  $|\mathbf{x}_0| \rightarrow \infty$

$$\lim_{|\mathbf{x}_0| \rightarrow \infty} \sqrt{\sum_{k=1}^n \sum_{i=1}^d a_i X_{k-i}} \left( \hat{\theta}_{|\mathbf{x}_0|,n} - \theta_0 \right) \stackrel{\mathcal{D}}{=} \mathcal{N}(0, \sigma^2(\theta_0))$$

$$\hat{\theta}_{|\mathbf{x}_0|,n} = 2.4486$$

## Paranoiac estimation

$$\hat{\theta}_{|\mathbf{x}_0|,n}^* = \arg \min_{\theta} \sum_{k=1}^n \frac{[x_k - \mathbb{E}_{\theta}(X_k^* | \mathbf{X}_{k-1}^* = \mathbf{x}_{k-1})]^2}{\sum_{i=1}^d a_i x_{k-i}}$$

- no explicit form of  $\hat{\theta}_{|\mathbf{x}_0|,n}^*$
- consistency as  $|\mathbf{x}_0| \rightarrow \infty$
- consistance et comportement asymptotique lorsque  $n \rightarrow \infty$

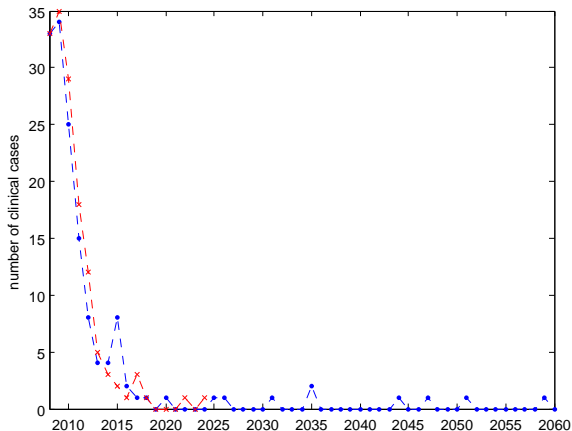
$$\lim_{n \rightarrow \infty} \sqrt{n} \left( \hat{\theta}_{|\mathbf{x}_0|,n}^* - \theta_0 \right) \stackrel{D}{=} \mathcal{N} \left( 0, \sigma^* (\theta_0)^2 \right)$$

$$\hat{\theta}_{|\mathbf{x}_0|,n}^* = 2.3977$$

$$\hat{\theta}_{|\mathbf{x}_0|,n}^* \leq \hat{\theta}_{|\mathbf{x}_0|,n}$$

# Simulation

Simulation of the *unconditioned* and *conditioned* process with infection parameter  $\theta_{env} = 2.3977$



# Prediction

Simulation of 1000 trajectories of the *conditioned* process with infection parameter  $\theta_{env} = 2.3977$

